

ECON 4151

Lab Session 4: Logit/Probit Model

Wonjin Lee



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Outline

1. Overview: LDV Models
2. Review of Probit/Logit Models (Wooldridge Ch. 17)

Limited Dependent Variable Models

- A limited dependent variable is a dependent variable whose range of values is substantively restricted.
- LDV Models:
 1. Binary dependent variable \implies logit and probit models
 2. Corner solution dependent variable \implies Tobit model
 - Lumpiness in an outcome variable (Ex. zero values for a nontrivial fraction of the population)
 3. Count dependent variable (nonnegative integer values) \implies Poisson regression models
 4. Data censoring (issue of data observability) \implies selectivity model

Binary response models

- Binary response models can be derived from an underlying **latent variable model**:

$$y_i^* = x_i' \beta + u_i$$

where u is independent of x .

- Observation rule for y : $y_i = \mathbb{1} [y_i^* > 0]$.
- Assume that u either has the **standard normal distribution** or the **standard logistic distribution**.

Binary response models

- We can derive the response probability for y :

$$\begin{aligned}\Pr[y_i = 1 \mid x_i] &= \Pr[y_i^* \geq 0 \mid x_i] \\ &= \Pr[u_i \geq -x_i'\beta] \\ &= 1 - \Pr[u_i \leq -x_i'\beta] \\ &= 1 - G(-x_i'\beta) \\ &= G(x_i'\beta) \quad (\because \text{symmetric distribution})\end{aligned}$$

where G is the cdf of u .

- In the probit model, G is the standard normal cdf.
- In the logit model, G is the logistic function.

Binary response models

- Partial effect (marginal effect):
 - For continuous explanatory variables,

$$\frac{\partial \Pr[y = 1|x]}{\partial x_j} = G'(x'\beta)\beta_j = g(x'\beta)\beta_j$$

- For discrete explanatory variables, the effect of x_k going from c_k to $c_k + 1$ is

$$G[\beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_k (c_k + 1)] \\ - G(\beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_k c_k)$$

- For logit and probit, the direction of the effect of x_j on the response probability and the latent variable is always the same.
- The magnitudes of each β_j are not useful because the latent variable rarely has a well-defined unit of measurement. Ex) the difference in utility levels from two different actions.

Binary response models

- Partial effect at the average:

$$\underbrace{g(\bar{x}'\hat{\beta})}_{\text{partial effect at the average}} \hat{\beta}_j$$

- Average partial effect:

$$\underbrace{\frac{1}{N} \sum_{i=1}^N g(x_i' \hat{\beta})}_{\text{average partial effect}} \hat{\beta}_j$$

Maximum likelihood estimation of binary response models

- In the linear probability model, the OLS estimator is unbiased but inefficient due to heteroskedasticity.
- MLE is based on the distribution of y given x , the heteroskedasticity in $\text{Var}(y|x)$ is automatically accounted for.

Maximum likelihood estimation of binary response models

- The density of y_i given x_i is

$$f(y_i | x_i' \beta) = G(x_i' \beta)^{y_i} (1 - G(x_i' \beta))^{1-y_i}, \quad y_i = 0, 1$$

- The joint distribution of the particular sequence we observe in a sample of N observations is simply:

$$f(y_1, y_2, \dots, y_N) = \prod_{i=1}^N G(x_i' \beta)^{y_i} (1 - G(x_i' \beta))^{1-y_i}$$

- We want to find $\hat{\beta}$ which maximizes the joint distribution; the joint distribution is a function of β

Maximum likelihood estimation of binary response models

- By taking the log of the joint distribution, we get the log-likelihood function

$$\log \mathcal{L}(\beta; y_1, y_2, \dots, y_N) = \sum_{i=1}^N [y_i \log G(x_i' \beta) + (1 - y_i) \log (1 - G(x_i' \beta))]]$$

- Find $\hat{\beta}$ s.t. $\frac{\partial \log \mathcal{L}}{\partial \beta} = 0$.
- Since $\frac{\partial \log \mathcal{L}}{\partial \beta}$ is nonlinear in β , we usually do not have explicit solution.
- As long as the model is correctly specified, MLE $\hat{\beta}$ is consistent, efficient, and asymptotically normal.
- $\hat{\beta}$ comes with an (asymptotic) standard error which can be used to construct (asymptotic) t tests and confidence intervals just as with OLS.

Excercise

- Linear probability model:

$$\begin{aligned} smoker = & \alpha_0 + \alpha_1 age + \alpha_2 educ \\ & + \alpha_4 income + \alpha_5 pcigs79 + \epsilon \end{aligned}$$

- Logit/probit model:

$$\begin{aligned} \Pr[smoker = 1|X] = & G(\beta_0 + \beta_1 age + \beta_2 educ \\ & + \beta_4 income + \beta_5 pcigs79) \end{aligned}$$

Probit Model

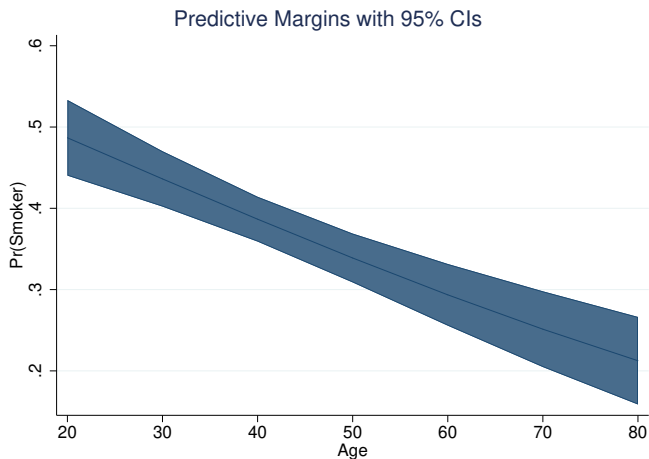


Figure: Predicted $\Pr[\text{smoke}]$ given age

Probit Model

- Examples from class:

$$\Pr[inlf = 1|X] = G(\beta_0 + \beta_1nwifeinc + \beta_2educ + \beta_3exper + \beta_5age + \beta_6kidslt6 + \beta_7kidsge6)$$

$$\Pr[inlf = 1|X] = G(\beta_0 + \beta_1nwifeinc + \beta_2educ + \beta_3exper + \beta_4exper^2 + \beta_5age + \beta_6kidslt6 + \beta_7kidsge6)$$

Probit Model

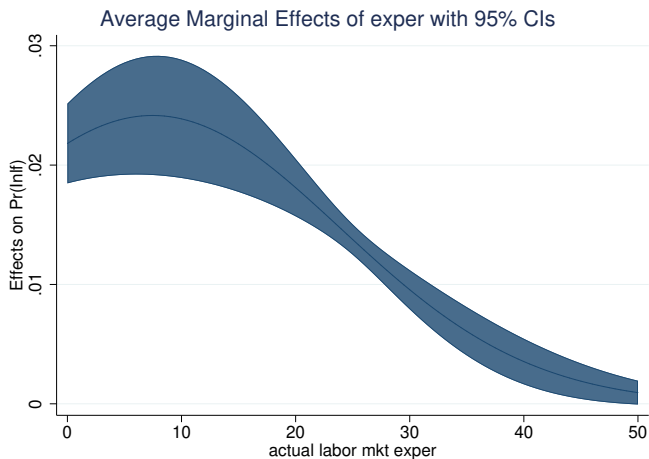


Figure: Average marginal effects of *exper* on $\Pr[\text{inlf}]$

Probit Model

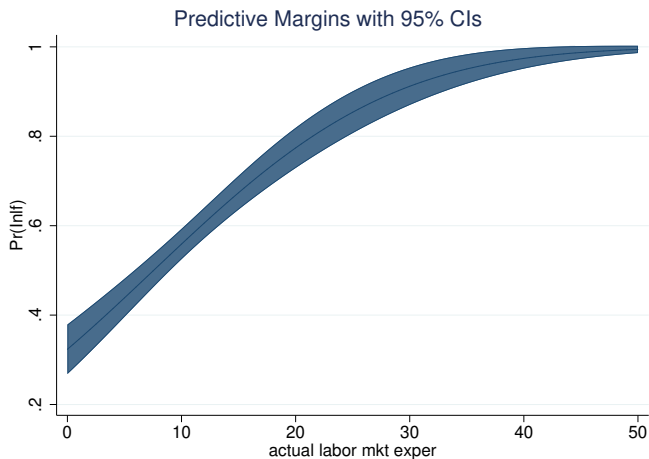


Figure: Predicted $\Pr[inlf]$ given $exper$